

Table of Discrete and Continuous distributions

$$E[X^k] = M^{(k)}(0)$$

$$\text{Var}[X] = E[X^2] - (E[X])^2 \\ = E[(X-\mu)^2]$$

$$M(t) = E[e^{tX}], M(0) = 1$$

Distribution	Type	Mass/density function $f(x)$	Mean μ	Variance σ^2	Moment Generating Function
UNIFORM(n)	D	$1/n$, for $x = 1, 2, \dots, n$	$(n+1)/2$	$(n^2-1)/12$	$\frac{e^{at} - e^{(b+1)t}}{(b-a+1)(1-e^t)}$, $t \in \mathbb{R}$
UNIFORM(a, b)	C	$\frac{1}{b-a}$, for $x \in [a, b]$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$	$\frac{e^{tb} - e^{ta}}{tb - ta}$, $t \in \mathbb{R}$
BERNOULLI(p)	D	$f(0) = 1-p, f(1) = p$	p	$p(1-p)$	$q + pe^t$, $t \in \mathbb{R}$
BINOMIAL(n, p)	D	$\binom{n}{x} p^x (1-p)^{n-x}$, for $x = 0, 1, \dots, n$	np	npq	$(q + pe^t)^n$, $t \in \mathbb{R}$
GEOMETRIC(p)	D	$q^{x-1}p$, for $x = 1, 2, \dots$	$1/p$	$(1-p)/p^2$	$\frac{p}{1-qe^t}$, $t < -\log q$
POISSON(λt)	D	$\frac{1}{x!} (\lambda t)^x e^{-\lambda t}$, for $x = 0, 1, \dots$	λt	λt	$e^{\lambda(t-1)}$, $t \in \mathbb{R}$
HYPERGEOMETRIC(N, M, n)	D	$\frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{n}}$, for $x = 0, 1, \dots, n$	np	$np(1-p)$	Too complex
EXPONENTIAL(λ)	C	$\lambda e^{-\lambda x}$, for $x \in [0, \infty)$	$1/\lambda$	$1/\lambda^2$	$\frac{\lambda}{\lambda-t}$, $t < \lambda$
GAMMA(α, β)	C	$\frac{x^{\alpha-1} e^{-x/\beta}}{\beta^\alpha \Gamma(\alpha)}$, for $x \in [0, \infty)$	$\alpha\beta$	$\alpha\beta^2$	$\left(\frac{1}{1-\beta t}\right)^\alpha$, $t < \frac{1}{\beta}$
NORMAL(μ, σ^2)	C	$\frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$, for $x \in \mathbb{R}$	μ	σ^2	$e^{\mu t + \frac{\sigma^2 t^2}{2}}$, $t \in \mathbb{R}$

$$N(0,1): E(\Sigma^n) = \begin{cases} 0 & n \text{ even} \\ (n-1)(n-3)\dots 3 \cdot 1 & n \text{ odd} \end{cases}$$

Degenerate (α)

$$D \quad P(\Sigma=x) = \begin{cases} 1 & x=a \\ 0 & x \neq a \end{cases}$$

a

0

$$e^{at}, t \in \mathbb{R}$$

Multinomial ($m, p_1, \dots, p_{K-1}$)

$$P(\Sigma_1=x_{11}, \dots, \Sigma_{K-1}=x_{K-1}) = \binom{m}{x_{11}, \dots, x_{K-1}} \prod_{j=1}^{K-1} p_j^{x_j}, \quad (x_{11}, \dots, x_{K-1}) = \frac{m!}{x_{11}! \dots x_{K-1}!}$$

$$\begin{cases} E\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i E[X_i] \\ \text{Var}\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i^2 \text{Var}(X_i) + \sum_{i,j} a_i a_j \text{Cov}(X_i, X_j) \end{cases}$$

$\bar{X}$ is IID: $\left\{ \begin{aligned} E\left(\sum_{i=1}^n X_i\right) &= n\mu \\ \text{Var}\left(\sum_{i=1}^n X_i\right) &= n\sigma^2 \end{aligned} \right.$

Σ_i IID: $M_{\Sigma_i, \mathbf{z}_i}(t) = \prod_{i=1}^n M_{\mathbf{z}_i}(t a_i)$

$$\left\{ \begin{array}{l} \Sigma_n \xrightarrow{P} \Sigma \Leftrightarrow \forall \epsilon > 0 : P(|\Sigma_n - \Sigma| > \epsilon) = 0 \text{ as } n \rightarrow \infty \\ \Sigma_n \xrightarrow{d} \Sigma \Leftrightarrow F_{\Sigma_n}(x) \rightarrow F_{\Sigma}(x) \quad \forall x \in \mathbb{R} \text{ as } n \rightarrow \infty \\ \quad \quad \quad \hookrightarrow \text{a continuity p of } F_{\Sigma} \end{array} \right.$$

$$\Sigma_n \xrightarrow{d} c \Rightarrow \overline{\Sigma}_n \xrightarrow{p} c$$

$$\Sigma_n \xrightarrow{p} \Sigma \Rightarrow \overline{\Sigma}_n \xrightarrow{d} \Sigma$$

WLOLN: (X_n) IID, $\begin{cases} E(X_i) = \mu < \infty \\ \text{Var}(X_i) = \sigma^2 < \infty \end{cases}$, $E(X_i^2) < \infty$ 8

$$\overline{X_n} = \frac{1}{n} \sum_{i=1}^n X_i \xrightarrow{P} \mu = E[X_i] \text{ as } n \rightarrow \infty$$

Stetsthe:

$$\left\{ \begin{array}{l} x_n \xrightarrow{d} x \\ y_n \xrightarrow{p} c \in \mathbb{R} \end{array} \right. \Rightarrow \left\{ \begin{array}{l} x_n + y_n \xrightarrow{d} x + c \\ x_n y_n \xrightarrow{d} cx \\ \underline{c \neq 0}: \frac{x_n}{y_n} \xrightarrow{d} \frac{x}{c} \end{array} \right.$$

Note : $C=0 \Rightarrow \sum n \gamma_n \frac{P_n d}{n} = 0$

CLT :

$$\left\{ \begin{array}{l} Z_n = \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \xrightarrow{d} Z \sim N(0,1) \\ Z_n = \frac{\sum_{i=1}^n X_i - n\mu}{\sqrt{n\sigma^2}} \xrightarrow{d} Z \sim N(0,1) \end{array} \right.$$

IID : $\begin{cases} E(X_i) = \mu < \infty \\ \text{Var}(X_i) = \sigma^2 < \infty \end{cases}$

$$\Rightarrow \begin{cases} \overline{X}_n \approx N(\mu, \frac{\sigma^2}{n}) \\ \sum_{i=1}^n X_i \approx N(n\mu, n\sigma^2) \end{cases} \text{ for } n \text{ large.}$$

$$E[g(X)] = \begin{cases} \sum_{x \in S_X} g(x) \cdot P(X=x) \\ \int_{-\infty}^{\infty} g(x) f_X(x) dx \end{cases}$$

$$E(|X|^t) < \infty \Rightarrow \begin{cases} E(|X|^s) < \infty \quad \forall s \leq t \\ \lim_{n \rightarrow \infty} n^t P(|X| > n) = 0 \end{cases}$$

$$\begin{cases} E[aX+b] = aE[X] + b \\ \text{Var}[aX+b] = a^2 \text{Var}[X] \end{cases}$$

$$p^{\text{th}} \text{ Quantile of } X: \begin{cases} P(X \leq x_p) \geq p \\ P(X \geq x_p) \leq 1-p \end{cases}, \text{ i.e. } p \leq F(x_p) \leq p + F(X=x_p)$$

$\rightarrow h(x) \geq 0 !!!$

$$E(h(X)) < \infty \Rightarrow P(h(X) \geq M) \leq \frac{E[h(X)]}{M}, \quad \forall M > 0$$

$$E(|X|^k) < \infty \Rightarrow P(|X| \geq M) \leq \frac{E[|X|^k]}{M^k}, \quad \forall M > 0 \quad (\text{Markov})$$

$$\left. \begin{matrix} E(X) = \mu, \text{Var}(X) = \sigma^2 \\ E(X^2) < \infty \end{matrix} \right\} \Rightarrow P(|X - \mu| \geq k\sigma) \leq \frac{1}{k^2}, \quad \forall k > 0 \quad (\text{Chebyshev-Bienayme})$$

$$E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i] \quad \text{if } E[|X_i|] < \infty$$

$$F(x, y) = \int_{-\infty}^x \int_{-\infty}^y f(x, y) dy dx \quad \Leftrightarrow \quad f(x, y) = \frac{\partial^2 F(x, y)}{\partial x \partial y}$$

$$\left. \begin{aligned} f_X(x) &= \int_{-\infty}^{\infty} f(x, y) dy & f_Y(y) &= \int_{-\infty}^{\infty} f(x, y) dx \\ F_X(x) &= \lim_{y \rightarrow \infty} F(x, y) & F_Y(y) &= \lim_{x \rightarrow \infty} F(x, y) \end{aligned} \right\} \begin{aligned} f_X(x) &= \frac{\partial F_X(x)}{\partial x} \\ f_Y(y) &= \frac{\partial F_Y(y)}{\partial y} \end{aligned}$$

$$\text{Cond PMF} = \frac{\text{Joint PMF}}{\text{Marginal PMF}} > 0 \quad ; \quad P(X=x_i | Y=y_j) = \frac{P_{ij}}{P_{\cdot j}} = \frac{P(X=x_i, Y=y_j)}{P(Y=y_j)}$$

$$\text{Cond PDF} = \frac{\text{Joint PDF}}{\text{Marginal PDF}} > 0 \quad ; \quad f_{X|Y=y}(x) = \frac{f(x, y)}{f_Y(y)}$$

$$\text{Cond CDF: } F_{X|Y=y}(x) = \int_{-\infty}^x f_{X|Y=y}(t) dt \quad ; \quad F_{X|Y=y_j}(x) = \sum_{k=1}^i \frac{P_{kj}}{P_{\cdot j}}$$

$$\begin{cases} E[X] = E[E[X|Y]] \\ E[g(X)] = E[E[g(X)|Y]] \end{cases}$$

$$\begin{cases} \text{Var}[X|Y=y] = E[(X - \mu_{X|Y=y})^2 | Y=y] = E[X^2|Y=y] - (E[X|Y=y])^2 \\ \text{Var}[X] = \text{Var}[E(X|Y)] + E[\text{Var}(X|Y)] \end{cases}$$

$$\text{Indep} = X \perp Y \Leftrightarrow F_{X,Y}(x,y) = F_X(x) F_Y(y) \Leftrightarrow \begin{cases} P(X=x, Y=y) = P(X=x) P(Y=y) \\ f_{X,Y}(x,y) = f_X(x) f_Y(y) \end{cases}$$

$$X \perp Y \Rightarrow \begin{cases} E(XY) = E(X)E(Y) \\ E(g_1(X)g_2(Y)) = E(g_1(X))E(g_2(Y)) \end{cases}$$

$$\text{Identically dist} = F_{X_i}(x) = F_{X_j}(x) \quad \forall i, j \in \{1, \dots, n\} \xrightarrow{\forall x \in \mathbb{R}} F_X(x) = F_Y(x) \quad \forall x \in \mathbb{R} \\ \hookrightarrow X_i \stackrel{d}{=} X_j$$

$$\text{Transfo} = \begin{cases} f_{Y_1, \dots, Y_n}(y_1, \dots, y_n) = f_{X_1, \dots, X_n}(g_1^{-1}(y_1, \dots, y_n), \dots, g_n^{-1}(y_1, \dots, y_n)) \cdot |J| \\ J = \det \begin{pmatrix} \frac{\partial x_1}{\partial y_1} & \dots & \frac{\partial x_1}{\partial y_n} \\ \vdots & \ddots & \vdots \\ \frac{\partial x_n}{\partial y_1} & \dots & \frac{\partial x_n}{\partial y_n} \end{pmatrix} \neq 0 \end{cases}$$

$$MGE = M_{\vec{X}}(t_1, \dots, t_n) = E[e^{t_1 X_1 + \dots + t_n X_n}] = \begin{cases} \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} e^{t_1 x_1 + \dots + t_n x_n} f(x_1, \dots, x_n) dx_1 \dots dx_n \\ \sum_{\vec{x} \in S_X} e^{t_1 x_1 + \dots + t_n x_n} P(X_1=x_1, \dots, X_n=x_n) \end{cases}$$

$$\hookrightarrow M_{\vec{X}}(0, \dots, 0) = 1, \quad M_{\vec{X}}(0, \dots, t_i, \dots, 0) = M_{X_i}(t_i)$$

$$\bullet E[X_1^m X_2^n] = \frac{\partial^{m+n} M_{\vec{X}}(t_1, t_2)}{\partial t_1^m \partial t_2^n}$$

$$\bullet X_i \perp \Rightarrow M_{\vec{X}}(t_1, \dots, t_n) = \prod_{i=1}^n M_{X_i}(t_i)$$

$$\bullet X_i \text{ IID} \Rightarrow M_{\vec{X}}(t_1, \dots, t_n) = \prod_{i=1}^n M_{X_1}(t_i) \rightsquigarrow M_{\vec{X}}(t_1, \dots, t_n) = [M_{X_1}(t_1)]^n$$

$$\text{Cauchy Schwarz} : 0 \leq E(|XY|) \leq \sqrt{E(X^2)E(Y^2)}$$

$$0 \leq E(|XY|) \leq (E(X^p))^{1/p} (E(Y^q))^{1/q}, \quad \frac{1}{p} + \frac{1}{q} = 1$$

$$\rightarrow \text{Cov}(X, Y) = \sigma_{X,Y} = E[(X - \mu_X)(Y - \mu_Y)] = E[XY] - E[X]E[Y] \quad \begin{cases} > 0 : \text{Concordant} \\ < 0 : \text{Discordant} \\ = 0 : \text{uncorr} \end{cases}$$

$$\bullet \text{Cov}(aX+b, cY+d) = ac \text{Cov}(X, Y)$$

$$\bullet \sigma_{XX} = \sigma_X^2$$

$$\hookrightarrow \text{indep} \Rightarrow \sigma_{X,Y} = 0$$

$$\rightarrow \text{Corr}(X, Y) = \rho_{X,Y} = \frac{\sigma_{X,Y}}{\sigma_X \cdot \sigma_Y} = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X) \text{Var}(Y)}}$$

$$\rightsquigarrow \text{measure linear rel}^{\circ} \text{ between } X, Y : \\ Y = aX + b \Leftrightarrow \rho_{X,Y} = \begin{cases} +1, & a > 0 \\ -1, & a < 0 \end{cases}$$

$$\bullet |\rho_{X,Y}| \leq 1, \quad \rho_{XX} = 1$$

$$\bullet \rho_{aX+b, cX+d} = \text{Sgn}(ac) \rho_{X,Y}$$